

# A Concentration of Measure & Random Matrix Perspective to Machine Learning

Random Matrix Seminar –  
Mathematical Institute, Oxford.  
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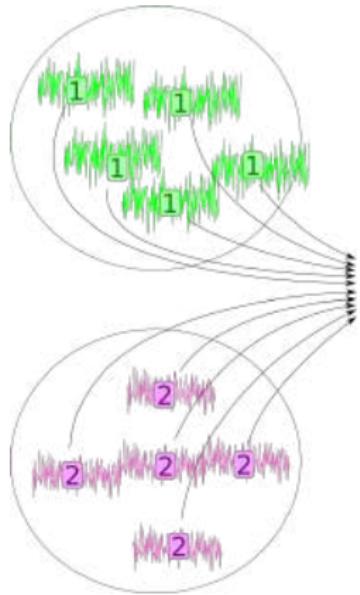
GIPSA-lab, INP Grenoble; List-CEA

18/02/2020

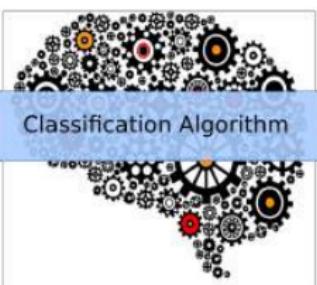
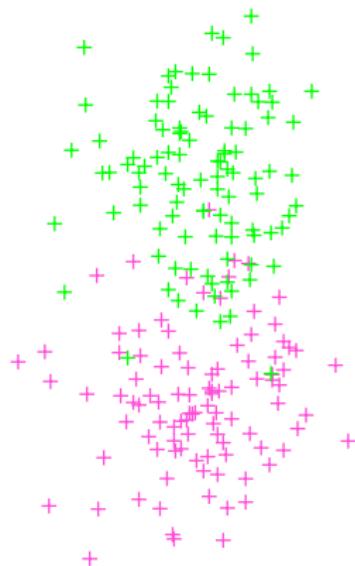


# Position of the problem

Gaussian data



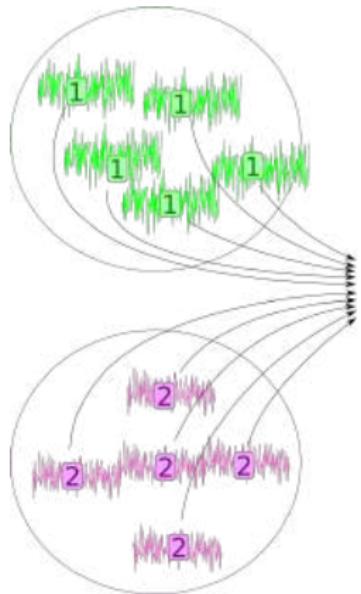
Classification output



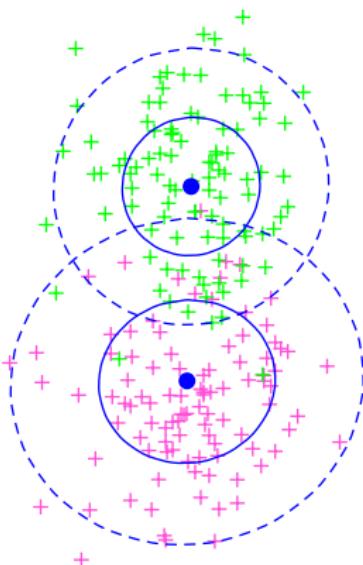
How to predict performances in high dimension ?

# Position of the problem

Gaussian data



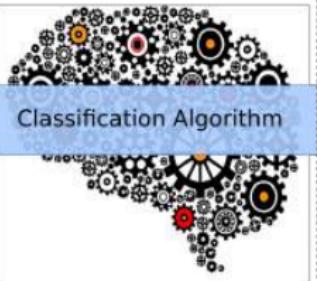
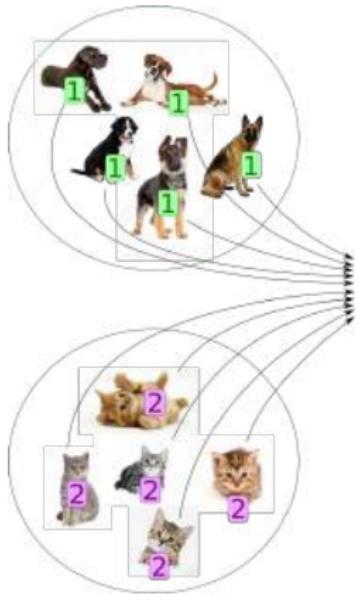
Predictions



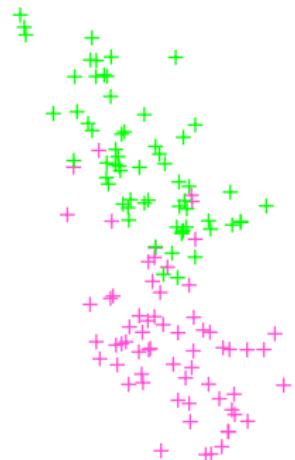
⇒ Resort to **Random Matrix Theory** tools

# Position of the problem

Real data



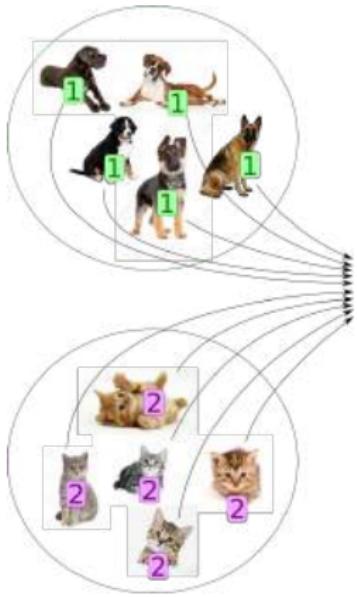
Classification output



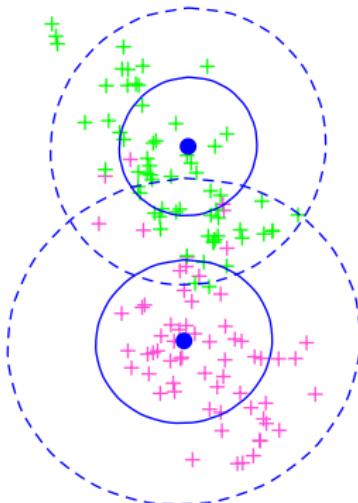
How to predict performances in realistic settings ?

# Position of the problem

Real data



Gaussian Predictions



⇒ Resort to RMT + Concentration of measure hypotheses

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- Design of the deterministic equivalent of Q
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# Classical study of singular values of rectangular RM

$X = (x_1, \dots, x_n) \in \mathcal{M}_{p,n}$ , spectral distribution of  $\frac{1}{n}XX^T$ :

$$\frac{1}{p} \sum_{\lambda \in \text{Sp}\left(\frac{1}{n}XX^T\right)} \delta_\lambda$$

## Classical Hypothesis

- ▶  $X$  has i.i.d entries with bounded variance
- ▶  $X = C^{\frac{1}{2}}Z$ ,  $Z \sim \mathcal{N}(0, I_n)$ .

## Classical conclusions

- ▶ Weak convergence of the spectral distribution to the Marcenko-Pastur law

**Question :** Can we find relaxed hypotheses and control the speed of convergence ?

# With the concentration of measure theory (CMT)

## Hypothesis of CMT

1. For all 1-Lipschitz maps  $f : \mathcal{M}_{p,n} \rightarrow \mathbb{R}$ :

$$\forall t > 0 : \mathbb{P}(|f(X) - \mathbb{E}[f(X)]| \geq t) \leq 2e^{-t^2/2}$$

(Independently on  $p$  and  $n$  !)

2. The column of  $X$  are i.i.d.

## Remarks

- ▶ **(Cons)** Implies all the moments are bounded
- ▶ **(Pros)** True if the columns are **Lipschitz transformations** of a Gaussian vector  $Z \sim \mathcal{N}(0, I_p)$ .  
→ dependence between entries of a column possibly complex

# With the concentration of measure theory (CMT)

## Conclusions on the spectral distribution

- ▶ Noting  $Q(z) = (\frac{1}{n}XX^T + zI_p)^{-1}$ , the resolvent of  $\frac{1}{n}XX^T$ ,  
 $(\frac{1}{p}\text{Tr}(Q(z)) : \text{Stieltjes transform})$

$\exists C, c \underset{p,n \rightarrow \infty}{=} O(1) :$

$$\forall t > 0 : \mathbb{P} \left( \left| \text{Tr}(AQ(z)) - \text{Tr}(A\tilde{Q}) \right| \geq t \right) \leq C \exp \left( - \frac{nt^2}{c\|A\|_1^2} \right)$$

where  $\tilde{Q} \in \mathcal{M}_p$  is a **deterministic equivalent** of  $Q$

(if  $A = \frac{1}{p}I_n$  : convergence of the Stieltjes transform of spectral dist. of  $\frac{1}{n}XX^T$ )

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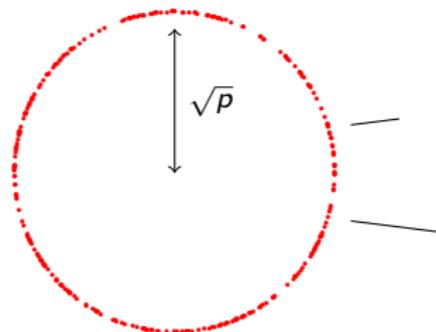
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# Concentration of Measure Phenomenon<sup>1</sup>

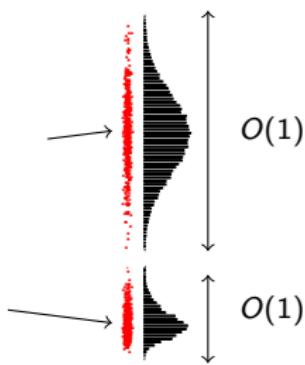
$$X = (X_1, \dots, X_p) \sim s_p$$



$$\frac{X_1 + \dots + X_p}{\sqrt{p}}$$

$$\|X\|_\infty$$

Observations



$$\begin{aligned} \text{Distribution diameter} &= \mathbb{E}[\|Z - \mathbb{E}Z\|] \\ &\underset{p \rightarrow \infty}{=} O(\sqrt{p}) \end{aligned}$$

$$\text{Observable diameter} \underset{p \rightarrow \infty}{=} O(1)$$

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<sup>1</sup>Ledoux - 2001 : The concentration of measure phenomenon

## Setting

$(E, \|\cdot\|)$ , a normed vector space,  $Z_{n,p} \in E$ , a random vector

- ▶  $(\mathbb{R}^p, \|\cdot\|)$ , with  $\|x\| = \sqrt{\sum_{i=1}^p x_i^2}$
- ▶  $(\mathcal{M}_{p,n}, \|\cdot\|_F)$  with  $\|M\|_F = \sqrt{\text{Tr}(MM^T)} = \sqrt{\sum_{\substack{1 \leq i \leq p \\ 1 \leq j \leq n}} M_{i,j}^2}$
- ▶  $(\mathcal{M}_{p,n}, \|\cdot\|)$  with  $\|M\| = \sup_{\|x\| \leq 1} \|Mx\|$

## Definition of concentration

if  $\exists C, c > 0$ ,  $(\sigma_{p,n})_{p,n \in \mathbb{N}} \in \mathbb{R}_+^{\mathbb{N}^2}$  |  $\forall n, p \in \mathbb{N}$ ,  
 $\forall f : (E, \|\cdot\|) \rightarrow (\mathbb{R}, |\cdot|)$  1-Lipschitz:

$$\boxed{\forall t > 0 : \mathbb{P}(|f(Z_{n,p}) - \mathbb{E}[f(Z_{n,p})]| \geq t) \leq Ce^{-(t/c\sigma_{p,n})^2}},$$

we note  $Z \propto \mathcal{E}_2(\sigma)$

## Fundamental example of the Theory:

$Z \in \mathbb{R}^p$ , if  $Z \sim \text{Unif}(\sqrt{p}\mathcal{S}^{p-1})$ ,  $Z \sim \text{Unif}(\mathcal{B}_{\mathbb{R}^p}(0, \sqrt{p}))$  or  
 $Z \sim \mathcal{N}(0, I_p)$ :

$\forall f : E \rightarrow \mathbb{R}$  1-Lipschitz :

$$\forall t > 0 : \mathbb{P}(|f(Z) - \mathbb{E}[f(Z')]| \geq t) \leq 2e^{-t^2/2},$$

Choosing  $C = 2, c = \sqrt{2}, \sigma_p = 1$  :

$$Z \propto \mathcal{E}_2(1) \text{ (Independent of } p \text{ !).}$$

→ Standard Hypothesis :  $Z \propto \mathcal{E}_2$

## Notion of deterministic equivalent

- if  $C, c > 0$ ,  $(\sigma_{p,n})_{p,n \in \mathbb{N}} \in \mathbb{R}_+^{\mathbb{N}^2}$  |  $\forall n, p \in \mathbb{N}$ ,  
 $\forall f : (E, \|\cdot\|) \rightarrow (\mathbb{R}, |\cdot|)$  1-Lipschitz :

$$\boxed{\forall t > 0 : \mathbb{P}(|f(Z) - \mathbb{E}[f(Z)]| \geq t) \leq Ce^{-(t/c\sigma_{p,n})^2}},$$

Notation:  $Z \propto \mathcal{E}_2(\sigma)$

- In particular, if  $\exists \tilde{Z} \in E$  |  $\forall u : E \rightarrow \mathbb{R}$  1-Lipschitz and linear :

$$\boxed{\forall t > 0 : \mathbb{P}\left(\left|u(Z - \tilde{Z})\right| \geq t\right) \leq Ce^{-(t/c\sigma_{p,n})^2}},$$

Notation:  $Z \in \tilde{Z} \pm \mathcal{E}_2(\sigma)$

$\tilde{Z}$  : Deterministic equivalent of  $Z$ .

Of course:  $Z \propto \mathcal{E}_2(\sigma) \implies Z \in \mathbb{E}[Z] \pm \mathcal{E}_2(\sigma)$

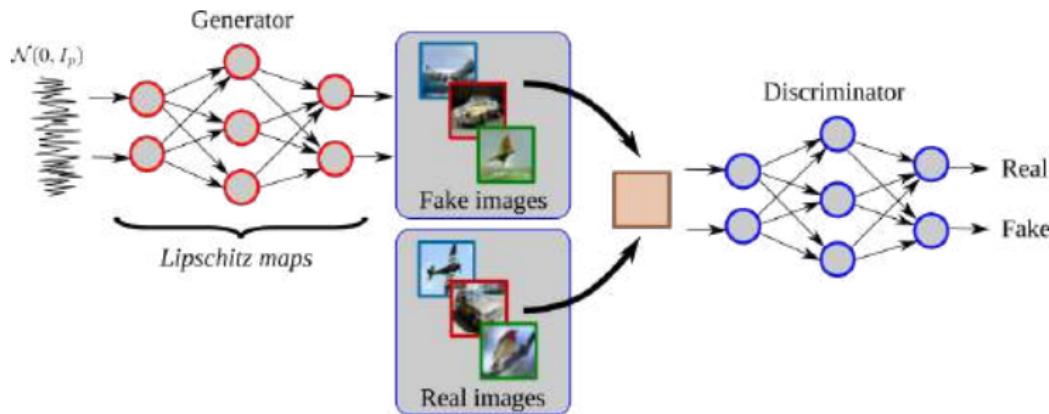
# Strategy for the study of $Q = (\frac{1}{n}XX^T + \gamma I_p)^{-1}$

1.  $Z \propto \mathcal{E}_2$  in  $(\mathbb{R}^p, \|\cdot\|)$   
 $\implies Q \propto \mathcal{E}_2(1/\sqrt{n})$  in  $(\mathcal{M}_{p,n}, \|\cdot\|_F)$   
 $\implies Q \in \mathbb{E}[Q] \pm \mathcal{E}_2(1/\sqrt{n})$  in  $(\mathcal{M}_{p,n}, \|\cdot\|_F)$
2.  $\exists \tilde{Q} \in \mathcal{M}_p$ ,  $\|\mathbb{E}[Q] - \tilde{Q}\| = O(1/\sqrt{n})$   
 $\implies Q \in \tilde{Q} \pm \mathcal{E}_2(1/\sqrt{n})$  in  $(\mathcal{M}_{p,n}, \|\cdot\|)$

# How to build new concentrated random vectors ?

- ▶ If  $Z \propto \mathcal{E}_2(\sigma)$  and  $f : E \rightarrow E$   $\lambda$ -Lipschitz,  $f(Z) \propto \mathcal{E}_2(\lambda\sigma)$
- ▶ No simple way to set the concentration of  $(Z_1, \dots, Z_p)$  if  $Z_1, \dots, Z_p \propto \mathcal{E}_2(\sigma)$ . Two possibilities:
  1.  $Z_1, Z_2$ , **independent** then:  $(Z_1, Z_2) \propto \mathcal{E}_2(\sigma)$
  2.  $(Z_1, Z_2) = f(Z)$  where  $Z \propto \mathcal{E}_2(\sigma)$ , and  $f$  1-Lipschitz.  
Then:  $(Z_1, Z_2) \propto \mathcal{E}_2(\sigma)$

# Realistic images built with GANs are concentrated



FAKE IMAGE =  $f(Z)$ , with  $f$  1 – Lipschitz and  $Z \sim \mathcal{N}(0, I_p)$



## Characterization with the moments

$$Z \propto \mathcal{E}_2(\sigma) \iff \begin{cases} \forall r \geq q, \forall f : E \rightarrow \mathbb{R}, \text{1-Lipschitz}, \exists c > 0 : \\ \mathbb{E}[|f(Z) - \mathbb{E}[f(Z)]|^r] \leq C \left(\frac{r}{q}\right)^{\frac{r}{q}} (c\sigma)^r \end{cases}$$

Proof ( $f \equiv f(Z)$  and  $\bar{f} = \mathbb{E}[f(Z)]$ ):

$\Rightarrow$  Fubini:

$$\begin{aligned} \mathbb{E}[|f - \bar{f}|^r] &= \int_Z \left( \int_0^\infty \mathbb{1}_{t \leq |f - \bar{f}|^r} dt \right) dZ \\ &= \int_0^\infty \mathbb{P}(|f - \bar{f}|^r \geq t) dt \\ &\leq \int_0^\infty C e^{-t^{\frac{q}{r}}/(c\sigma)^q} dt \dots \leq C' \left(\frac{r}{q}\right)^{\frac{r}{q}} \sigma^r \end{aligned}$$

$\Leftarrow$  Markov inequality:

$$\mathbb{P}(|f - \bar{f}| \geq t) \leq \frac{\mathbb{E}[|f - \bar{f}|^r]}{t^r} \leq C \left(\frac{r}{q}\right)^{\frac{r}{q}} \left(\frac{c\sigma}{t}\right)^r,$$

$$\text{with } r = \frac{qt^q}{e(c\sigma)^q} \geq q : \mathbb{P}(|f - \bar{f}| \geq t) \leq C e^{-(t/c\sigma)^q/e}.$$



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## Key result : Control of the norm

- ▶ Infinite norm :

$$\begin{aligned}\mathbb{P} \left( \|Z - \tilde{Z}\|_\infty \geq t \right) &= \mathbb{P} \left( \sup_{1 \leq i \leq p} e_i^T (Z - \tilde{Z}) \geq t \right) \\ &\leq p \sup_{1 \leq i \leq p} \mathbb{P} \left( e_i^T (Z - \tilde{Z}) \geq t \right) \\ &\leq C e^{\log p - (t/c\sigma)^2} \leq C' e^{-(t/\sigma\sqrt{\log(p)})^2},\end{aligned}$$

- ▶ For the general case, use of “ $\varepsilon$ -nets”. If  $\exists H \subset (E^*, \|\cdot\|_*)$  |

$$\forall z \in E : \|z\| = \sup_{f \in H} f(z).$$

$$Z \in \tilde{Z} \pm C\mathcal{E}_q(\sigma) \implies \|Z - \tilde{Z}\| \in 0 \pm 8^{\dim(\text{Vect}(H))} C\mathcal{E}_q(2\sigma)$$

on  $(\mathbb{R}^p, \|\cdot\|)$ ,  $H = \mathbb{R}^p$ , and  $\dim(\text{Vect}(H)) = p$



## Norm degree

Degree of a subset  $H \subset E^*$  and of a norm

- ▶  $\eta_H = \log(\#H)$  if  $H$  is finite
- ▶  $\eta_H = \dim(\text{Vect}(H))$  if  $H$  is infinite

Degree of a norm

- ▶  $\eta_{\|\cdot\|} = \inf \left\{ \eta_H, H \subset E^* \mid \forall x \in E, \|x\| = \sup_{f \in H} f(x) \right\}$

Example

- ▶  $\eta(\mathbb{R}^p, \|\cdot\|_\infty) = \log(p)$
- ▶  $\eta(\mathcal{M}_{p,n}, \|\cdot\|) = n + p$
- ▶  $\eta(\mathbb{R}^p, \|\cdot\|) = p$
- ▶  $\eta(\mathcal{M}_{p,n}, \|\cdot\|_F) = np.$

## Concentration of the norm

If  $Z \in \tilde{Z} \pm C\mathcal{E}_q(\sigma)$ :

$$\|Z - \tilde{Z}\| \in 0 \pm \mathcal{E}_q(\sigma \sqrt{\eta_{\|\cdot\|}}) \quad \text{and} \quad \mathbb{E} \|Z - \tilde{Z}\| = O\left(\sigma \sqrt{\eta_{\|\cdot\|}}\right).$$

Example  $Z \in \mathbb{R}^p$ ,  $X \in \mathcal{M}_{p,n}$

- ▶ if  $Z \in \tilde{Z} \pm \mathcal{E}_2$  :  $\mathbb{E} \|Z\| \leq \|\tilde{Z}\| + C\sqrt{p}$
- ▶ if  $X \in \tilde{X} \pm \mathcal{E}_2$  :  $\mathbb{E} \|X\| \leq \|\tilde{X}\| + C\sqrt{p+n}$ ,
- ▶ if  $X \in \tilde{X} \pm \mathcal{E}_2$  :  $\mathbb{E} \|X\|_F \leq \|\tilde{X}\|_F + C\sqrt{pn}$ .

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## Concentration of the sum and the product

If  $(X, Y) \in \mathcal{E}_2(\sigma)$  ( $X, Y$  independent or  $(X, Y) = f(Z)$ ,  $Z \in \mathcal{E}_2(\sigma)$ ):

- ▶  $X + Y \in \mathcal{E}_2(\sigma)$
- ▶  $(X - \tilde{X})(Y - \tilde{Y})$

$$\propto \mathcal{E}_1(\sigma^2) + \mathcal{E}_2\left(\sigma^2 \sqrt{\eta_{\|\cdot\|'}}\right) \text{ in } (E, \|\cdot\|)$$

where  $\forall x, y \in \mathcal{E}$   $\|xy\| \leq \|x\|'\|y\|$  (usually  $\|x\|' \leq \|x\|$ ).

Example  $X \in \mathcal{M}_{p,n}$ ,  $Y, Z \in \mathbb{R}^p$ ,  $Y, Z, X \propto \mathcal{E}_2$

- ▶  $\frac{XX^T}{n} \propto \mathcal{E}_2\left(\frac{\sqrt{n+p}}{n}\right) + \mathcal{E}_1\left(\frac{1}{n}\right)$  in  $(\mathcal{M}_{p,n}, \|\cdot\|_F)$
- ▶  $Y \odot Z \propto \mathcal{E}_2(\sqrt{\log p}) + \mathcal{E}_1$  in  $(\mathbb{R}^p, \|\cdot\|)$

## Hanson Wright-like results

### Classical Theorem<sup>2</sup>

If  $Z_1, \dots, Z_p \in \mathbb{R}$ , independent,  $\forall i, Z_i \sim \mathcal{E}_2$ ,  $\mathbb{E}[Z_i] = 0$ :

$$\pi \equiv \mathbb{P} \left( |Z^T A Z - \mathbb{E} Z^T A Z| \geq t \right) \leq C \exp \left( -c \min \left( \left( \frac{t}{\|A\|_F} \right)^2, \frac{t}{\|A\|} \right) \right)$$

With the Concentration of the measure phenomenon

If  $Z = (Z_1, \dots, Z_p) \sim \mathcal{E}_2$ ,  $\mathbb{E}[Z_i] = 0$  two results:

$$1. \pi \leq C \exp \left( -c \min \left( \left( \frac{t}{\sqrt{p}\|A\|} \right)^2, \frac{t}{\|A\|} \right) \right)$$

$$2. \pi \leq C \exp \left( -c \min \left( \left( \frac{t}{\sqrt{\log p}\|A\|} \right)^2, \frac{t}{\|A\|_F} \right) \right)$$

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<sup>2</sup>Roman Vershynin - High-Dimensional Probability

## Proofs

$A = S_+ - S_- + R$  with  $S_+, S_- \geq 0$  and  $R$  antisymmetric  
 $\Rightarrow$  enough to prove the result for  $A \geq 0$

1.  $Z^T AZ \propto \mathcal{E}_2(\sqrt{p} \|A\|) + \mathcal{E}_1(\|A\|)$  :

$Z^T AZ = \|A^{1/2}Z\|^2$  where  $A^{1/2}Z \propto \mathcal{E}_2(\|A\|^{1/2})$  and  
 $\mathbb{E}[\|A^{1/2}Z\|] = \sqrt{n}\|A\|^{1/2}$ .

2.  $Z^T AZ \propto \mathcal{E}_2(\sqrt{\log p} \|A\|_F) + \mathcal{E}_1(\|A\|_F)$

$A = P^{-1}\Lambda P$ , with  $\Lambda = \text{Diag}(\lambda_1, \dots, \lambda_p)$

$Y \equiv PZ \propto \mathcal{E}_2$  (since  $\|P\| = 1$ ) and:

$$Z^T AZ = (Y \odot Y)^T \lambda$$

But  $Y \odot Y \propto \mathcal{E}_2(\sqrt{\log p}) \pm \mathcal{E}_1$   
 $\rightarrow$  we conclude since  $\|\lambda\| = \|A\|_F$

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## Position of the problem

Data matrix  $X = (x_1, \dots, x_n) \in \mathcal{M}_{p,n}$ ,

### Hypothesis:

- ▶  $p = O(n)$  and  $n = O(p)$
- ▶  $X \propto \mathcal{E}_2$
- ▶  $\|\mathbb{E}[X]\| = O(\sqrt{n})$

### Strategy:

1. concentration of the resolvent:  $Q = Q(z) = \left(\frac{1}{n}XX^T + zI_p\right)^{-1}$ ,
2. computable *deterministic equivalent*  $\tilde{Q}$ ,
3. spectral dist. of  $\frac{1}{n}XX^T$  from estimation of Stieltjes transf:

$$m(z) = \frac{1}{p} \operatorname{Tr}(Q(z)).$$

$$1 - \text{Concentration of } Q = Q(z) = \left( \frac{1}{n} X X^T + z I_p \right)^{-1}$$

$Q = f(X)$  with  $f$   $O(\frac{1}{\sqrt{n}})$ -Lipschitz and  $X \propto \mathcal{E}_2$  thus:

$$Q \in \mathbb{E}[Q] \pm \mathcal{E}_2 \left( \frac{1}{\sqrt{n}} \right)$$

## 2 - Choice of a **computable** deterministic equivalent

- ▶ naive choice :  $\tilde{Q} = (\Sigma + zI_p)^{-1} \rightarrow \text{wrong!}$
- ▶ clever choice :  $\tilde{Q} = \left( \frac{\Sigma}{1+\delta} + zI_p \right)^{-1}$  with
  1.  $\delta = \frac{1}{n} \text{Tr}(\Sigma \mathbb{E}[Q]) \rightarrow \text{not directly computable !}$
  2.  $\delta$  solution to:

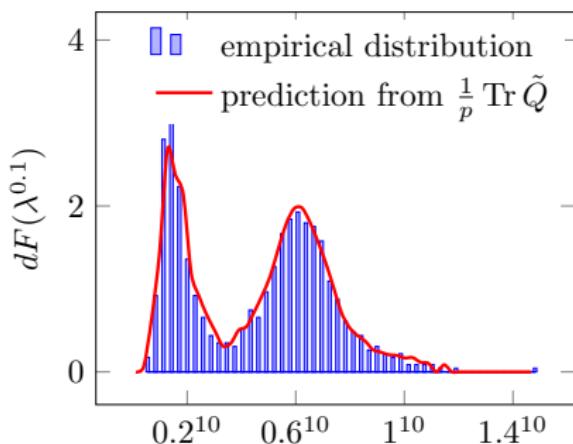
$$\delta = \frac{1}{n} \text{Tr} \left( \Sigma \left( \frac{\Sigma}{1+\delta} + zI_p \right)^{-1} \right)$$

# Conclusion for the spectral distribution of $\frac{1}{n}XX^T$

## Theorem

$Q \in \tilde{Q} \pm \mathcal{E}_2\left(\frac{1}{\sqrt{n}}\right)$  in  $(\mathcal{M}_{p,n}, \|\cdot\|)$  in particular,  $\forall z > 1$ :

$$\mathbb{P}\left(\left|m(z) - \frac{1}{p} \operatorname{Tr}(\tilde{Q}(z))\right| \geq t\right) \leq Ce^{-(nt/c)^2}, \text{ for } C, c = O(1)$$



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# Regression in Machine Learning

## Setting with 2 classes

- ▶ 2 laws in  $\mathbb{R}^p$  :  $\mathcal{C}_+$ ;  $\mathcal{C}_-$
  - ▶  $X = (x_1, \dots, x_n) \in \mathcal{M}_{p,n}$ : data matrix,  $x_i \sim \mathcal{C}_+$  or  $x_i \sim \mathcal{C}_-$
  - ▶ notation:  $\mu_a = \mathbb{E}[x_i]$ ,  $\Sigma_a = \mathbb{E}[x_i x_i^T]$ , for  $x_i \sim \mathcal{C}_a$
  - ▶  $Y \in \{-1, 1\}^n$ : label vector  $x_i \sim \mathcal{C}_a \Rightarrow y_i = a$
- Look for  $\beta \in \mathbb{R}^p$  s.t.  $X^T \beta \approx Y$ .

## Ridge Regression

Minimise  $\frac{1}{n} \| \beta^T X - Y \|^2 + \gamma \| \beta \|^2$ ,  $\gamma$  : regularising parameter

## Robust Regression<sup>3</sup>

Minimise  $\frac{1}{n} \sum_{i=1}^n f(y_i \beta^T x_i) + \gamma \| \beta \|^2$ , for a loss function  $f : \mathbb{R} \rightarrow \mathbb{R}$ .

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<sup>3</sup>El Karoui - 2013 : On robust regression with high-dimensional predictors

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# Ridge Regression

$$\text{Minimise} \quad \frac{1}{n} \sum_{i=1}^n (\beta^T x_i - y_i)^2 + \gamma \|\beta\|^2.$$

**Solution :**  $\beta = \frac{1}{n} QXY$  with  $Q = \left(\frac{1}{n} XX^T + \gamma I_p\right)^{-1}$ .

## Performance estimation

► **Training error:**  $E_{\text{tr}} = \frac{1}{n} \|X^T \beta - Y\|^2$

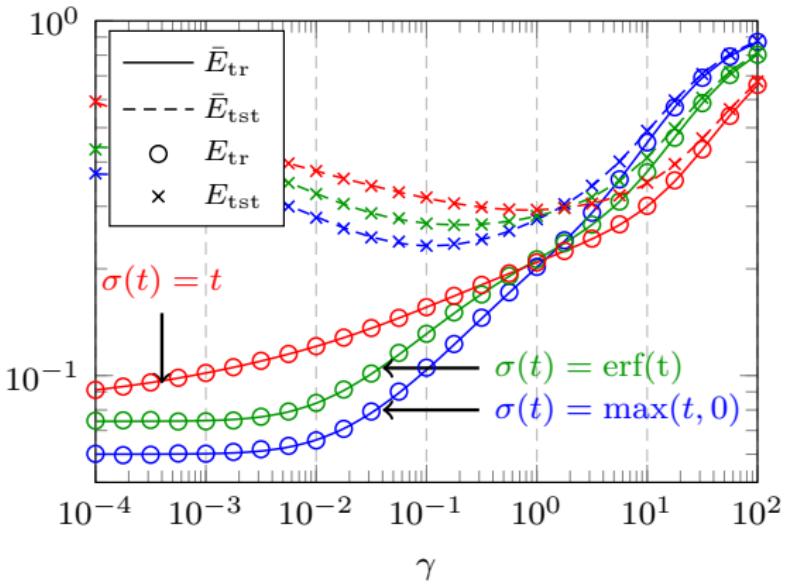
$$\bar{E}_{\text{tr}} = \frac{1}{n} \mathbb{E} \left[ \left\| \frac{1}{n} X^T QXY - Y \right\|^2 \right] = f_1^\circ(\Sigma_\pm, \mu_\pm).$$

► **Test error:**  $E_{\text{tst}} = \frac{1}{n} \|X_t^T \beta - Y\|^2$ ,  $X_t, X$  i.i.d

$$\bar{E}_{\text{tst}} = \frac{1}{n} \mathbb{E} \left[ \frac{1}{n} Y X Q X_t X_t^T Q X Y - 2 Y^T X_t^T Q X Y + Y^T Y \right] = f_2^\circ(\Sigma_\pm, \mu_\pm)$$

# Example with One-Layer Neural Net $X = \sigma(WZ)$

- ▶  $Z = (z_1, \dots, z_n) \in \mathcal{M}_{q,n}$ , MNIST data
- ▶  $W \in \mathcal{M}_{p,q}$ , fixed initial drawing
- ▶  $\sigma : \mathbb{R} \rightarrow \mathbb{R}$ : Lipschitz activation function



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# Robust Regression

$$\text{Minimise} \quad \frac{1}{n} \sum_{i=1}^n f(y_i x_i^T \beta) + \gamma \|\beta\|^2$$

**Solution :**  $\beta = \frac{1}{n\gamma} \sum_{i=1}^n \phi(z_i^T \beta) z_i$  with  $z_i = y_i x_i$  and  $\phi = -\frac{1}{2} f'$

## Theorem

If  $X \in \mathcal{E}_2$ ,  $\phi : \mathbb{R} \rightarrow \mathbb{R}$  is  $\lambda$ -Lipschitz bounded and  $\gamma > \frac{1}{\sqrt{n}} \lambda \|\mathbb{E}[X]\|^2$   
then  $\beta$  is uniquely defined and:

$$\beta \in \mathcal{E}_2 \left( \frac{1}{\sqrt{n}} \right) \quad \text{and} \quad \mathbb{E}[\|\beta\|] = O(1).$$

→ estimation of  $\mathbb{E}[\beta]$  and  $\mathbb{E}[\beta \beta^T]$  to predict performances



## Statistics of $\beta$

### Definition of $\beta_{-i}(t)$

$$\text{Sol}^\circ \text{ of } \beta_{-i}(t) = \frac{1}{n\gamma} \sum_{\substack{j=1 \\ j \neq i}}^n \phi(z_j^T \beta_{-i}(t)) z_j + \frac{1}{n} \underbrace{t \phi(z_i^T \beta_{-i}(t))}_{\chi_i(t)} z_i.$$

$\beta_{-i} = \beta_{-i}(0)$  and  $\beta = \beta_{-i}(1)$ .

### 1 - differentiation of $\beta_{-i}(t)$

$$\begin{aligned} \beta'_{-i}(t) &= \frac{1}{n\gamma} \sum_{\substack{j=1 \\ j \neq i}}^n \underbrace{\phi'(z_j^T \beta_{-i}(t)) z_j z_j^T}_{D_i(t)} \beta'_{-i}(t) + \frac{1}{n} \chi'_i(t) z_i \\ &= \frac{1}{n\gamma} Z_{-i} D(t) Z^T \beta'_{-i}(t) + \frac{1}{n} \chi'_i(t) z_i, \end{aligned}$$

$$\text{with } Q(t) = \left( \frac{1}{n} Z_{-i} D(t) Z^T - \frac{1}{\gamma} I_p \right)^{-1} : \beta'_{-i}(t) = \frac{1}{n} \chi'_i(t) Q(t) z_i$$

# Statistics of $\beta$

2 - Link between  $z_i^T \beta_{-i}$  and  $z_i^T \beta$

$$z_i^T \beta'_{-i}(t) = \chi'(t) \frac{1}{n} z_i^T Q(t) z_i$$

with  $Q(t) = \left( \frac{1}{n} Z_{-i} D(t) Z_{-i}^T - \frac{1}{\gamma} I_p \right)^{-1}$  and  $\chi(t) = t \phi(z_i^T \beta_{-i}(t))$

## Proposition

$\left| \frac{1}{n} z_i^T Q(t) z_i - \frac{1}{n} z_i^T Q(0) z_i \right| = O(\frac{1}{\sqrt{n}})$  and with  $\delta \equiv \mathbb{E}[\frac{1}{n} z_i^T Q(0) z_i]$ :

$$\frac{1}{n} z_i^T Q(t) z_i \in \delta \pm \mathcal{E}_2 \left( \frac{1}{\sqrt{n}} \right) + \mathcal{E}_1 \left( \frac{1}{n} \right)$$

## Integration

$$\begin{aligned} z_i^T \beta - z_i^T \beta_{-i} &= \delta(\chi(1) - \chi(0)) + O\left(\frac{1}{\sqrt{n}}\right) \\ &= \delta \phi(z_i \beta) + O\left(\frac{1}{\sqrt{n}}\right) \end{aligned}$$



# Statistics of $\beta$

## 3 - computation of $\mathbb{E}[\beta]$ and $\mathbb{E}[\beta\beta^T]$

- ▶  $\beta_{-i}$  and  $z_i$  independent  $\Rightarrow z_i^T \beta_{-i} \sim \mathcal{N}(m_z^T m_\beta, \text{Tr}(C_z C_\beta))$
- ▶ Deduce<sup>4</sup>  $m_\beta$  and  $\Sigma_\beta$  from:
  - ▶  $z_i^T \beta - z_i^T \beta_{-i} \approx \delta \phi(z_i^T \beta)$
  - ▶  $\beta = \frac{1}{n} \sum_{i=1}^n \phi(z_i^T \beta) z_i$

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<sup>4</sup>Mai, Liao, Couillet - A Large Scale Analysis Of Logistic Regression: Asymptotic Performances And New Insights

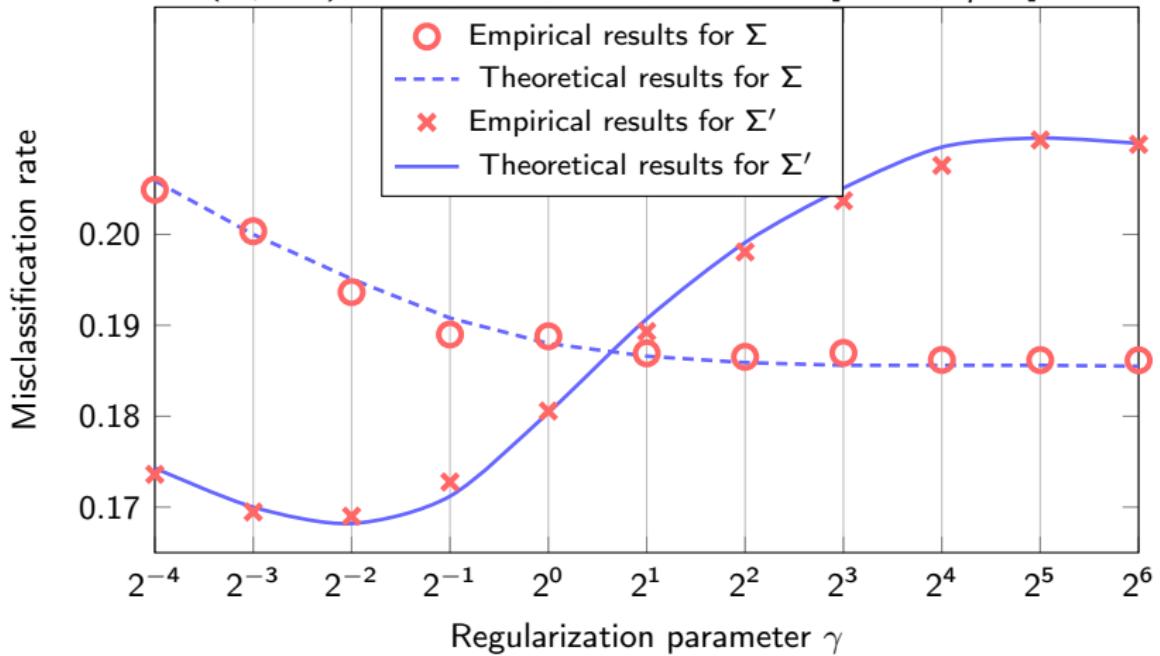
# Application

►  $p = 128, n = 512$

►  $x_i \sim \mathcal{N}(y_i\mu, \Sigma)$

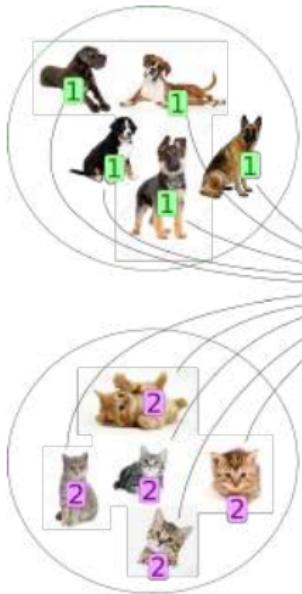
►  $\Sigma = 2I_p$

►  $\Sigma' = \text{diag}[1, 5, \mathbf{1}_{p-2}]$



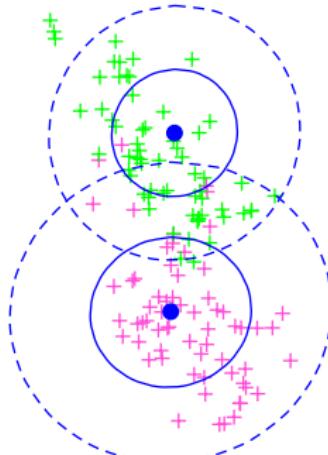
# Conclusion

Real images  
~ Concentrated vectors



Ridge Regression,  
One-layer NN,  
Robust Regression...

RMT + CMT tools  
for Perf. Predictions



# Conclusion

Random matrix theory allows for:

- ▶ precise estimate when  $n = O(p)$

Combination with Concentration of Measure allows for

- ▶ extension to **realistic** hypotheses (GAN-generated data),
- ▶ tracking of concentration through **explicit** and **implicit** formulations,
- ▶ **rich** and **adaptable** characterisation of relevant quantities.

Thank you !

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## Position of the problem

Data matrix  $X = (x_1, \dots, x_n) \in \mathcal{M}_{p,n}$ ,

Hypothesis:

- ▶  $p = O(n)$  and  $n = O(p)$
- ▶  $X \in \mathcal{E}_2(c)$
- ▶  $\|\mathbb{E}[X]\| = O(\sqrt{n})$

Goal:

Show the concentration of the resolvent:

$$Q = Q(z) = \left( \frac{1}{n} X X^T + z I_p \right)^{-1}$$

and find a computable *deterministic equivalent*  $\tilde{Q}_1$  depending on the population covariance :  $\Sigma = \frac{1}{n} \mathbb{E}[X X^T]$

# Basic results on the resolvent $Q = \left(\frac{1}{n}XX^T + zI_p\right)^{-1}$

- The resolvent is **bounded**:

$$\|Q(z)\| \leq \frac{1}{z}, \quad \left\| Q(z) \frac{XX^T}{n} \right\| \leq 1 \text{ and } \left\| Q(z) \frac{X}{\sqrt{n}} \right\| \leq \frac{1}{z^{1/2}}$$

- $X \mapsto Q(z)$  is  $\frac{1}{\sqrt{n}z^{3/2}}$ -**Lipschitz**:

If we note  $Q(z)^H = \left(\frac{1}{n}(X + H)(X + H)^T + zI_p\right)^{-1}$  :

$$\begin{aligned}\left\| Q(z)^H - Q(z) \right\|_F &= \left\| \frac{1}{n} Q(z)^H (XX^T - (X + H)(X + H)^T) Q(z) \right\|_F \\ &= \left\| -\frac{1}{n} Q(z)^H H X^T + (X + H) H^T Q(z) \right\|_F \\ &\leq \frac{1}{\sqrt{n}} \left( \|Q(z)^H\| \left\| \frac{1}{\sqrt{n}} X^T Q \right\| + \left\| \frac{1}{\sqrt{n}} Q^H (X + H) \right\| \|Q(z)\| \right) \|H\|_F\end{aligned}$$

- $Q(z) \in \mathbb{E}[Q(z)] \pm C\mathcal{E}_2 \left( \frac{c}{\sqrt{n}} \right)$  (we suppose that  $\frac{1}{z} = O(1)$ )

## Question

How to estimate  $\mathbb{E} \left[ \left( \frac{1}{n} \mathbf{X} \mathbf{X}^T + z I_p \right)^{-1} \right]$  ?

## Design of a Deterministic equivalent

Let  $\tilde{\Sigma} \in \mathcal{M}_p$  to be chosen precisely later and we set:

$$\tilde{Q}_1 = (\tilde{\Sigma} + z I_p)^{-1}$$

With identity  $A^{-1} - B^{-1} = A^{-1}(B - A)B^{-1}$ :

$$\mathbb{E}[\tilde{Q}_1 - Q] = \mathbb{E}\left[Q\left(\frac{1}{n}XX^T - \tilde{\Sigma}\right)\tilde{Q}_1\right] = \sum_{i=1}^n \frac{1}{n}\mathbb{E}\left[Q(x_i x_i^T - \tilde{\Sigma})\tilde{Q}_1\right].$$

## Schur formulas

We set  $X_{-i} = (x_1, \dots, x_{i-1}, 0, x_{i+1}, \dots, x_n) \in \mathcal{M}_{p,n}$  and  $Q_{-i} = (\frac{1}{n}X_{-i}X_{-i}^T + zI_p)^{-1}$ :

$$Q = Q_{-i} - \frac{1}{n} \frac{Q_{-i}x_i x_i^T Q_{-i}}{1 + \frac{1}{n}x_i^T Q_{-i}x_i} \quad \text{and} \quad Qx_i = \frac{Q_{-i}x_i}{1 + \frac{1}{n}x_i^T Q_{-i}x_i}.$$

Then:

$$\begin{aligned}\tilde{Q}_1 - \mathbb{E}Q &= \sum_{i=1}^n \frac{1}{n}\mathbb{E}\left[Q_{-i}\left(\frac{x_i x_i^T}{1 + \frac{1}{n}x_i^T Q_{-i}x_i} - \tilde{\Sigma}\right)\tilde{Q}_1\right] \\ &\quad - \frac{1}{n^2} \sum_{i=1}^n \mathbb{E}\left[Q_{-i}x_i x_i^T Q \tilde{\Sigma} \tilde{Q}_1\right].\end{aligned}$$



## A first deterministic equivalent

$$\begin{aligned}\left\| \tilde{Q}_1 - \mathbb{E}Q \right\| &= \sup_{\|u\|, \|v\| \leq 1} u^T (\tilde{Q}_1 - \mathbb{E}Q) v \\ &= \sup_{\|u\|, \|v\| \leq 1} \frac{1}{n} \sum_{i=1}^n \Delta_i + \varepsilon_i\end{aligned}$$

with:

- ▶  $\Delta_i = \mathbb{E} \left[ u^T Q_{-i} \left( \frac{x_i x_i^T}{1 + \frac{1}{n} x_i^T Q_{-i} x_i} - \tilde{\Sigma} \right) \tilde{Q}_1 v \right]$
- ▶  $\varepsilon_i = \frac{1}{n} \mathbb{E} \left[ u^T Q_{-i} x_i x_i^T Q \tilde{\Sigma} \tilde{Q}_1 v \right]$

→ we note  $\delta_1 = \frac{1}{n} \text{Tr}(\Sigma \mathbb{E}[Q_{-i}])$  and we chose  $\tilde{\Sigma} = \frac{\Sigma}{1 + \delta_1}$

Let us show that with this choice:  $\Delta_i, \varepsilon_i = O\left(\frac{1}{\sqrt{n}}\right)$



## Preliminary lemmas

- ▶  $u^T Q x_i = \frac{1}{\sqrt{n}} (\sqrt{n} u^T Q) x_i \in \mathcal{E}_2(c) + C \mathcal{E}_1 \left( \frac{c}{\sqrt{p}} \right)$
- ▶  $\mathbb{E}[u^T Q x_i] \leq \sqrt{\mathbb{E}[u^T Q x_i x_i^T Q u]} = \sqrt{\frac{1}{n} \mathbb{E}[u^T Q X X^T Q u]}.$   
 $\leq \mathbb{E} [\|u^T Q u\|] = O(1)$
- ▶ The same way:  
 $u^T Q_{-i} x_i, u^T \tilde{Q}_1 x_i \in O(1) \pm C \mathcal{E}_2(c) + C \mathcal{E}_1 \left( \frac{c}{\sqrt{p}} \right)$

## Preliminary lemmas

- ▶  $\frac{1}{n} \mathbf{x}_i^T Q_{-i} \mathbf{x}_i \in \mathcal{E}_2(c) + C\mathcal{E}_1\left(\frac{c}{\sqrt{n}}\right)$
- ▶  $\mathbb{E} \left[ \frac{1}{n} \mathbf{x}_i^T Q_{-i} \mathbf{x}_i \right] = \frac{1}{n} \text{Tr}(\Sigma \mathbb{E}[Q_{-i}]) \leq \frac{1}{n} \text{Tr}(\Sigma) \mathbb{E} [\|Q_{-i}\|] = O(1)$
- ▶
$$\begin{aligned} \|\mathbb{E} Q_{-i} - \mathbb{E} Q\| &= \sup_{\|u\|, \|v\| \leq 1} u^T (\mathbb{E} Q_{-i} - \mathbb{E} Q) v \\ &= \sup_{\|u\|, \|v\| \leq 1} \mathbb{E} \left[ \frac{1}{n} u^T Q_{-i} \mathbf{x}_i \mathbf{x}_i^T Q v \right] = \sup_{\|u\|, \|v\| \leq 1} \frac{1}{n} \sqrt{\mathbb{E} [ } \end{aligned}$$

End of the proof,  $\tilde{Q}_1 = \left( \frac{\Sigma}{1+\delta_1} + zI_p \right)^{-1}$

- Since  $\|\tilde{\Sigma}\tilde{Q}_1\| = O(1)$ , with Holder inequality :

$$\begin{aligned}\varepsilon_i &= \frac{1}{n} \mathbb{E} \left[ u^T Q_{-i} x_i x_i^T Q \tilde{\Sigma} \tilde{Q}_1 v \right] \\ &\leq \frac{1}{n} \sqrt{\mathbb{E} [(u^T Q_{-i} x_i)^2] \mathbb{E} [(x_i^T Q \tilde{\Sigma} \tilde{Q}_1 v)^2]} = O\left(\frac{1}{n}\right)\end{aligned}$$

$$\begin{aligned}\Delta_i &= \mathbb{E} \left[ u^T Q_{-i} \left( \frac{x_i x_i^T}{1 + \frac{1}{n} x_i^T Q_{-i} x_i} - \frac{\Sigma}{1 + \delta_1} \right) \tilde{Q}_1 v \right] \\ &= \mathbb{E} \left[ \frac{u^T Q_{-i} x_i x_i^T \tilde{Q}_1 v (\delta_1 - \frac{1}{n} x_i^T Q_{-i} x_i)}{(1 + \frac{1}{n} x_i^T Q_{-i} x_i) (1 + \delta_1)} \right] \\ &\quad + \mathbb{E} \left[ u^T Q_{-i} \left( \frac{x_i x_i^T - \Sigma}{1 + \delta_1} \right) \tilde{Q}_1 v \right] = O\left(\frac{1}{\sqrt{n}}\right)\end{aligned}$$

$$\implies \|\mathbb{E}[Q] - \tilde{Q}_1\| = O\left(\frac{1}{\sqrt{n}}\right)$$



## Second deterministic equivalent

Note that  $\delta_1 = \frac{1}{n} \text{Tr}(\Sigma \mathbb{E}[Q]) = \frac{1}{n} \text{Tr}(\Sigma \tilde{Q}_1) + O\left(\frac{1}{\sqrt{n}}\right)$

$$= \frac{1}{n} \text{Tr}\left(\Sigma \left(\frac{\Sigma}{1+\delta_1} + zI_p\right)^{-1}\right) + O\left(\frac{1}{\sqrt{n}}\right)$$

The function

$$\begin{aligned} \mathbb{R}^+ &\longrightarrow \mathbb{R}^+ \\ \delta &\longmapsto \frac{1}{n} \text{Tr}\left(\Sigma \left(\frac{\Sigma}{1+\delta} + zI_p\right)^{-1}\right) \end{aligned}$$

is contracting for the semimetric:  $d_s(\delta, \delta') = \frac{|\delta - \delta'|}{\sqrt{\delta \delta'}}$   
⇒ It admits a unique fixed point:

$$\delta_2 = \frac{1}{n} \text{Tr}\left(\Sigma \left(\frac{\Sigma}{1+\delta_2} + zI_p\right)^{-1}\right)$$



## End of the proof

It can be showed that  $\delta_1 - \delta_2 = O\left(\frac{1}{\sqrt{n}}\right)$  thus if we set

$$\tilde{Q}_2 = \left( \frac{\Sigma}{1+\delta_2} + zI_p \right)^{-1}:$$

$$\begin{aligned}\|\mathbb{E}[Q] - \tilde{Q}_2\| &\leq \|\mathbb{E}[Q] - \tilde{Q}_1\| + \|\tilde{Q}_1 - \tilde{Q}_2\| \\ &\leq O\left(\frac{1}{\sqrt{n}}\right) + \left\| \tilde{Q}_1 \frac{\Sigma(\delta_2 - \delta_1)}{(1 + \delta_2)(1 + \delta_1)} \tilde{Q}_2 \right\| = O\left(\frac{1}{\sqrt{n}}\right)\end{aligned}$$

$$\begin{aligned}&\Rightarrow Q \in \tilde{Q}_1 \pm C\mathcal{E}_2\left(\frac{c}{\sqrt{n}}\right) \\ &\Rightarrow \forall t > 0 : \mathbb{P}\left(\left|\frac{1}{p} \text{Tr}(AQ) - \frac{1}{p} \text{Tr}(A\tilde{Q}_2)\right| \geq t\right) \leq Ce^{-cnt^2},\end{aligned}$$

(for  $A \in \mathcal{M}_p$ ,  $\|A\|_1 \leq p$ )



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